

Some Special Cases of Khintchine's Conjectures in Statistical Mechanics:

Approximate Ergodicity of the Auto-Correlation Functions

of an Assembly of Linearly Coupled Oscillators

by Joseph F. Johnson

Dept. of Mathematics and Statistics, Villanova University

Abstract

In order to estimate the theoretical auto-correlation function of a time series from the sample auto-correlation function of one of its realisations, it is usually assumed without justification that the time series is ergodic. In 1943, Khintchine made some visionary conjectures about dynamical systems with large numbers of degrees of freedom which would justify, even in the absence of ergodicity, approximately the same conclusions. We prove Khintchine's conjectures in some special cases of a linearly coupled assembly of harmonic oscillators.

KEY WORDS: Time series, ergodic, auto-correlation function, statistical mechanics.

RESUMEN

Para emplear el correlograma de los valores muestrales de un proceso estocástico para estimar su función teórica de autocorrelación, por regla general se asume, sin justificación, que el proceso es ergódico. Pero en 1943, Khintchine conjeturó proposiciones de gran importancia en este asunto, que justificarían una aproximación a las mismas estimaciones aún sin la ergodicidad del sistema. Mostraremos casos particulares de las conjeturas de Khintchine para ensamblajes de osciladores lineales.

Preface

A novel way to justify the use, in Statistical Mechanics, of the equality of time averages with phase averages was envisioned by Khintchine in 1943, but he could only prove special cases [9], [10]. He suggested that for quite general, non-ergodic dynamical systems, a kind of approximate ergodicity for a restricted class of observables should arise when the number of degrees of freedom is sufficiently large. ([9], pp. 62–63).

In the study of Brownian motion by Ford–Kac–Mazur in 1965 [2] in terms of a Hamil-

tonian heat bath they carried out the Gibbs program in detail for a concrete, linear, Hamiltonian model.

The task of this paper is to recast the model of Brownian motion to show that one may assume a determined initial condition and derive a stochastic process in the thermodynamic limit without assuming any initial probability distribution. It extends Khintchine's vision to an essentially new case (earlier extensions by Ruelle [13] and Lanford [8] assumed weak, short-range interactions). It is already known that the results of [2] remain true for many different choices of an initial distribution (see Kim [11] for a survey).

The systems we study are linear Hamiltonian systems and are very far from being ergodic. But as Khintchine foresaw, a kind of approximate ergodicity holds good for some measurable functions (having particular physical significance) when the number of degrees of freedom is very large. In our case, the measurable functions we study are the auto-correlations of the time-evolution of the momentum coordinate of one of the particles.

This paper falls into two halves which are almost separate: the first half is a survey of the problem and does not pretend to any originality, given at a seminar at the Univ. of Havana in 2011. The second half presents the technical details which were left out of the seminar and are original.

Introduction to the Role of Ergodicity in the Theory of Time Series

Time Series: Two Contradictory Definitions

The notion of time series has two definitions which although related cause confusion to the student. The first meaning is that a time series is a series of data distributed in time. If, for example, M is a dynamical system with Hamiltonian $H(p, q) = \frac{1}{2}(|p|^2 + |q|^2)$, where p and q are $n + 1$ -dimensional vectors, i.e., a collection of $n + 1$ uncoupled harmonic

oscillators, then since $\dot{p} = -\frac{\partial H}{\partial q}$ and $\dot{q} = \frac{\partial H}{\partial p}$, $p(t)$ is a time series if $p(0)$ and $q(0)$ are given.

Another example is the (infinite) sequence

$$0111101111111111110010111111111101111111 \dots$$

of tosses of a fair coin (there would not be any contradiction of either the laws of Physics or the laws of Probability if all future tosses resulted in 1). These are both examples of deterministic data since even the coin toss is a function of time in the usual, deterministic, sense of the word *function*.

The second meaning of *time series*, is that it is a sequence of random variables. This sense is also called a *stochastic process*. Two examples which are related to the previous examples are:

Coin Toss: X_n a sequence of independent, identically distributed random variables taking the values 1 and -1 with equal probabilities. In fact, the space of all possible sequences of results of a coin toss, i.e., the space of all possible sequences of binary digits, can be mapped to the unit interval $[0, 1] \subset \mathbb{R}$ by regarding each sequence x_n as the binary expansion of the real number

$$\sum_0^{\infty} \frac{(1+x_n)}{2^{n+2}}.$$

This map is an equivalence of probability spaces (it is one to one except on a set of measure zero), between the space of all possible sequences of tosses and the unit interval with Lebesgue measure.

Dynamical System: put some (any) probability distribution on the set of initial conditions $\{p(0), q(0)\} = \mathbb{R}^{2(n+1)}$, for example, the Maxwell–Boltzmann distribution

$$(2\pi kT)^{n+1} \det A^{-\frac{1}{2}} e^{\frac{-H(p(0), q(0))}{kT}}$$

where k is Boltzmann's constant and T is the absolute temperature in Kelvin. (For later generalisations, we here allow the possibility of a linear coupling given by a matrix A .) Then $p(0)$ is now a random variable and so is $p(t)$ for any t , so $\{p_o(t)\}_t$ is a continuous series of random variables and hence a time series in the sense of a stochastic process.

From the standpoint of the rest of statistics, time series are odd and difficult because we have to regard the time series in the second sense as the population and the time series in the first sense as one sample taken from the population. For example, if the probability space of the random variables X_n from the example of the coin toss is taken to be the unit interval $[0, 1]$ with Lebesgue measure dx , then for any fixed $\alpha \in [0, 1]$ we get a time series of data, that is, a time series in the first sense, given by $X_n(\alpha)$. (Every sample point has probability zero, so this is an example of a probability space where probability zero does not mean 'impossible'.) Unlike the examples of data in first-year statistics courses, we can never draw more than one sample point since, e.g., we cannot go back to the year 2000 and 'try again'.

Other terms used are, e.g., that $\{p_o(t)\}_t$ is an ensemble of time series (in the first sense), and that given a particular value $(p(0), q(0))$ for the initial conditions, then p_o is a well-defined function of t called a *realisation* of the time series. These two senses are intimately related, the confusion in terminology serves a useful purpose, and it is not going to be reformed any time soon.

Statistics of Time Series.

The usual descriptive statistics from first-year statistics courses are less useful here. The average is misleading if the time series possess a trend. Trends and cycles are more important than the average or dispersion. The most important descriptive statistic of a

time series is its auto-correlation function or correlogram. (I wish to emphasise here that it is a descriptive statistic: it has no more probabilistic significance than did the average or standard deviation.) Given a series of data $f(t)$ it measures the average influence of $f(t)$ on $f(t + \tau)$ and is given by

$$\varphi(\tau) = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T f(t)f(t + \tau)dt.$$

This is called the sample auto-correlation function when it is necessary to distinguish it from a related notion which does not use the time average but uses the whole theoretical model (population) instead of only one realisation of the process (the data), the phase or population or model or ‘theoretical’ auto-correlation function:

$$R(\tau) = E(X_n X_{n+\tau}).$$

Here, E , the expectation, is taken over the probability space, which is sometimes the phase space of a dynamical system, so it can be regarded as a phase average as opposed to the sample auto-correlation function, which was a time average of actual given data. (For simplicity we assume, from now on, that all random variables are *centred*, i.e., have zero expectation.)

Examples.

1. Given a data stream $p(t) = \sum a_n \cos nt + \sum b_n \sin nt$, we obtain

$$\varphi(\tau) = \sum \frac{|a_n|^2 + |b_n|^2}{2} \cos n\tau,$$

which is an even function.

2. If the process $X(t) = p(t)$ ‘with random phases’, i.e., if each $p(t)$ is turned into a random variable by introducing random phase shifts in its terms, then $R(\tau)$ is the same thing. (The phases can be uniformly distributed or Gaussian, it makes no difference.)

3. Coin toss:

$$E(X_n X_m) = 0 \text{ if } n \neq m$$

by independence. $R(\tau) = \delta(\tau)$, a spike.

4. Data from a coin toss: $\varphi_\alpha(\tau)$ could be any positive definite function (that takes the value 1 at the origin). For example, one realisation could have a periodic correlogram, a sawtooth alternating between 1 and -1, and another set of tosses (all heads) could yield a constant function. Neither of these is very close to the theoretical auto-correlation function calculated above, but a ‘normal’ realisation will have a correlogram close to a spike.

Since we only have one sample point from the population, the problem is how to infer $R(\tau)$ from $\varphi(\tau)$? This is the topic of this paper. The answer has usually been taken to be the concept of *ergodicity*, a concept borrowed from Statistical Mechanics to which we will turn in the next section. The reader should be warned that within the discipline of time series, the term ‘large sample theory’ has been perverted from its meaning in the rest of statistics, since here its original meaning is largely irrelevant (and we will not use it in this lecture). Within advanced time series texts, it means the theory of one sample point which has a lot of data in it. (See the careful discussion in Fuller [3], pp. 308ff.)

Lévy’s philosophy.

Lévy pioneered the method of replacing the study of a stochastic process by a study of its auto-correlation function (sometimes called the auto-covariance function or sometimes normalised in a certain fashion). His philosophy [12] was that for a wide class of stochastic processes, all important properties can be seen in the theoretical auto-correlation function of the process.

Statistical Mechanics

If M is a dynamical system, as above, and one supposes that even non-engineering data sets such as those of climate change or coin tosses are indeed the results of an immensely complicated dynamical system with an astronomical number of degrees of freedom, then a measurable function f on M is called an observable. But in fact f itself is not observable. A measurement of f is always *macroscopic*, it is always the result of letting some part of the system come into contact with a measurement apparatus, such as a thermometer, and it takes time for the apparatus, which is of macroscopic dimensions, to react to the system and reach an equilibrium state. No state which changes rapidly, at a molecular scale, can be observed by the human eye, so we always model a measurement as an infinite time average and define the following notation:

$$\langle f \rangle_t = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T f(p(t), q(t)) dt.$$

The point is that $(p(0), q(0))$ itself is unknown and uncontrollable. Hence time averages are impossible to calculate, and yet they are what can be measured scientifically.

On the other hand, if we take $dpdq$ to be Liouville measure on M , then phase averages (we introduce two different notations for this same concept in the following equation)

$$\overline{f} = \langle f \rangle = \int_{\Omega} f(p, q) d\omega$$

can be calculated, at least approximately. Here, Ω is a compact surface of constant energy within M and the measure is the appropriate invariant measure inherited from Liouville measure.

A dynamical system is said to be ergodic if for all measurable f , we have

$$\langle f \rangle_t = \langle f \rangle$$

for almost all initial conditions $(p(0), q(0))$ (note that the left hand side depends implicitly on a choice of initial conditions but the right hand side depends only on f).

The importance of ergodicity is that if a dynamical system is ergodic, then macroscopic measurements, the only ones we can make, are reliable guides to the phase averages, the only quantities we can really calculate. Without ergodicity, there is no way to connect theory with experiment.

If a time series is ergodic then we can use the sample mean to estimate the mean, and also use the correlogram to estimate the theoretical auto-correlation function, which then by Lévy's philosophy tells us everything of interest about the stochastic process.

Linear systems are the opposite of ergodic. In fact, very few physical systems are known to be ergodic. In the 60's, Sinai proved that a system of an ideal billiard ball was ergodic. In 1941, Oxtoby and Ulam proved that 'most' dynamical systems are ergodic. Nevertheless, there is no proof that, e.g., the dynamical system of the weather or coin tossing is ergodic.

Khintchine [9] in 1943 proved that if $R(\tau) \rightarrow 0$ as $\tau \rightarrow \infty$, then f is ergodic, meaning that the above equation holds for almost all initial conditions. But R is not φ , and in particular it can not be observed directly and it depends on the choice of μ the probability measure. For linear systems, R is quasi-periodic and so is φ .

Statistical Mechanics considers dynamical systems in which the number of degrees of freedom is very large. Ocean waves can either be modelled by a non-linear wave equation such as Navier–Stokes, with a small number of degrees of freedom, or by a linear Hamiltonian mechanics, at the molecular level, with an astronomical number of degrees of

freedom. Even better, it considers a family of ‘similar’ dynamical systems parametrised by the number of degrees of freedom and considers various limits as $n \rightarrow \infty$. These limits include the traditional ‘thermodynamic limit’ but as Balian and others have argued, one can define many different types of such limits.

Khintchine observed that since the number of degrees of freedom is very large, asymptotic formulae for the quantities of interest should be obtainable by means of the methods of probability theory, especially its limit theorems. In particular, he observed that ergodicity in itself was not of central importance since it was asking for too much to have exact equalities of time averages and phase averages for all measurable functions. It would suffice to have relations for some physically significant observables which hold asymptotically as the number of degrees of freedom goes to infinity. In 1943 he published vague but profound and visionary conjectures in this regard, but was unable to establish them in more than a few special cases and even then only with the help of the assumption that $0 = 1$, for which he has been much criticised.

Khintchine’s conjectures

For a family of dynamical systems M_n , parametrised by their (increasing) number of degrees of freedom, and representing in some sense ‘the same physics’, and for certain physically significant quantities, each one represented by an f_n , an observable for each M_n , again in some sense being ‘the same’ as n increases, Khintchine conjectured that f_n would become approximately ergodic for n sufficiently large. Ruelle [13] and Lanford [8] were able to make some progress on this for systems with weak and short-range interactions and for observables that were some sort of average over the entire system, similar

to the thermodynamic quantities such as temperature. Yet Brownian motion is a very well-known ergodic stochastic process which does not at all fit into this framework: the momentum of one particle becomes, as the number of other particles increases without bound, a stationary stochastic process closely related to Brownian motion, known as the Ornstein–Uhlenbeck process. Since it is the momentum of only one particle, it is not a thermodynamic quantity nor do the methods of Khintchine–Ruelle–Lanford apply.

The Gibbs Program and Brownian Motion

In 1965 Ford–Kac–Mazur [2] showed how Brownian motion could arise in the limit of a sequence of explicit Hamiltonian systems. Their procedure was a model of carrying out the program envisioned by Willard Gibbs as long ago as 1900. The breakthrough was to allow a very violent, long-range interaction between the particles, one so violent as to require a kind of renormalisation in the limit. (In 1961 Schwinger [14] published a very interesting quantum precursor of this. Indeed, Schwinger’s set-up involved a negative temperature amplifier which amplified quantum motion into a classical stochastic process.) This successfully carried out Gibbs’s program for statistical mechanics for this concrete example, certainly one of central importance, and all the more striking since each system was linear but the limit stochastic process was ergodic. But they imposed a probability distribution (as did Gibbs himself, and in this respect was criticised by Khintchine), that of Maxwell–Boltzmann, on the dynamical systems by *fiat*, and did not address Khintchine’s conjectures.

Their results ought to be robust in the choice of probability distribution and the choice of interaction. Students of Kac, Kim [11], and others have pursued this question of robustness.

The key result of Ford–Kac–Mazur is as follows: fix a temperature T . Put the corresponding Maxwell–Boltzmann probability distribution on the space of initial conditions \mathbb{R}^{2n+2} . Then there exists a family of matrices A_n , each one giving a linearly coupled system of harmonic oscillators with Hamiltonian H_n , such that, with the appropriate cut-off in the interaction to avoid singularities,

$$R(\tau) \longrightarrow e^{-d|\tau|}$$

as $n \rightarrow \infty$ where R is the theoretical (phase) auto-correlation function of p_o , the momentum of the zeroth particle (both of which depend on n), and d is a constant. Since one knows the limit of the theoretical auto-correlation functions, then, by Lévy’s philosophy, one knows which stochastic process (up to equivalence) ought to be considered the limit of these processes. This is all the more striking since the coupling constants, the entries of the matrices A_n , do not possess a limit, but instead grow without bound.

Yet for any finite n , $p_o(t)$ is quasi-periodic:

$$p_o(t) = \sum a_n \cos \lambda_{n,k} t + \sum b_n \sin \lambda_{n,k} t$$

where the $\lambda_{n,k}$ depend on A_n . Hence so is R_n and so is φ_n . Hence $R_n \not\rightarrow 0$ as $\tau \rightarrow \infty$ for any n . We cannot interchange the limits in n and τ .

A sequence of dynamical systems

For future use, we introduce some common notation. For f any function on the phase space Ω of a dynamical system, let f_t denote the function composed with the flow on the system for t units of time. Let $\langle f \rangle = \int_{\Omega} f(\omega) d\mu$ whatever the invariant measure $d\mu$. Let $\langle f \rangle_t = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T f_t(\omega) dt$ which implicitly depends on $\omega \in \Omega$ although this will usually

be suppressed in the notation. The point is to investigate when $\langle f \rangle_t = \langle f \rangle$ approximately for almost all ω or at least the overwhelming majority of ω .

Ford–Kac–Mazur introduced a stylised model [2] of Brownian motion, which consists of a mote whose canonical co-ordinates are p_o, q_o , and n particles of the same mass. We are going to assume $n = 2N$ is even and let all vector and matrix indices run from $-N$ to N . The canonical co-ordinates of the i^{th} harmonic oscillator are p_i, q_i . Let the Hamiltonian of this system be H_n and write

$$H_n = \sum_{i=-N}^N \frac{p_i^2}{2m} + \frac{1}{2}(q_{-N}, q_{-N+1}, \dots, q_N)A \begin{pmatrix} q_{-N} \\ q_{-N+1} \\ \vdots \\ q_N \end{pmatrix}$$

where A is a symmetric $n + 1$ -square real matrix with positive eigenvalues.

The trajectories of this flow satisfy (where A depend implicitly on n .)

$$p(t) = \cos(A^{1/2}t) \cdot p(0) - A^{1/2} \sin(A^{1/2}t) \cdot q(0).$$

We focus on $p_0(t)$. If the particles are all alike, it is natural to assume the matrix A is what is called, ‘cyclic’. Each row is the previous one shifted over by one. The eigenvalues of A , ω_i^2 , satisfy

$$(A)_{ml} = \frac{1}{n+1} \sum_{-N}^N \omega_i^2 e^{\frac{2\pi\sqrt{-1}}{n+1}i(m-l)}$$

where $i \neq \sqrt{-1}$. This is obviously symmetric if we make a simple assumption on the ω_i^2 ’s.

Let $\zeta = e^{\frac{i\pi}{2N+1}}$. It is a classical fact about cyclic matrices that

$$(\cos A^{\frac{1}{2}}\tau)_{mn} = \frac{1}{n+1} \sum_{i=-N}^N \cos(\omega_i\tau) \zeta^{i(m-n)},$$

This formula is an expression of the fact that the vectors (there are $2N + 1$ of them as i runs from $-N$ to N) $\frac{1}{\sqrt{2N+1}}(\zeta^{ij})_{j=-N, \dots, N}$ are a normal basis of eigenvectors of A with eigenvalues ω_i^2 .

This formula holds for $A^{\frac{1}{2}} \sin A^{\frac{1}{2}} t$ as well. (We omit the proofs of facts about cyclic matrices, which may be found in Gerhard Kowalewski, *Determinantentheorie* (as cited in Ford–Kac–Mazur [2])). Hence (here and elsewhere, all indices run from $-N$ to N)

$$p_o(t) = \frac{1}{n+1} \left\{ \sum_k \sum_i \cos(\omega_i t) \zeta^{-ik} p_k(0) - \sum_k \sum_i \omega_i \sin(\omega_i t) \zeta^{-ik} q_k(0) \right\}.$$

We put $\hat{p}(k) = \sum_i \zeta^{-ik} p_i(0)$ and similarly for \hat{q} and rewrite the sums above as

$$p_o(t) = \frac{1}{n+1} \left\{ \sum_k \hat{p}(k) \cos(\omega_k t) - \sum_k \hat{q}(k) \omega_k \sin(\omega_k t) \right\}.$$

Define the auto-correlation (sometimes called the auto-covariance) function of this trajectory by $\varphi_n(\tau) = \langle p_o(t) p_o(t + \tau) \rangle_t$. For each τ , $\varphi_n(\tau)$ is a physical observable on M_n . (As τ varies, we have a uniform family of physical observables.) Just as all the M_n have, intuitively, the same physical meaning, so too these observables (or, rather, uniform families of observables) all have the ‘same physical meaning’. Intuitively, each one measures how ‘random’ the trajectory (through a given point of phase space) is. It is a descriptive statistic of a definite set of data: whatever the initial conditions ‘really are’, the data is the future path of the trajectory (or the past, it makes no difference), and this is a deterministic descriptive statistic. The main result of this paper is to make rigorous the notion that in the limit as $n \rightarrow \infty$, all normal trajectories have the same autocorrelation function, the Markoffian one which represents maximal randomness possible in this situation.

But since this intuitive notion of limit is problematic, we make this notion rigorous by talking about normal cells for finite n . For fixed n , we can define a normal trajectory as being one which (within certain limits of approximation) has the same auto-correlation function as all other normal trajectories, viz., the one which is the best possible approximation to the Markoffian exponential decay.

In general, the physicist Sir James Jeans outlined a common-sensical view of the foundations of Statistical Mechanics which has had an influence on his followers, Darwin and Fowler, and through them, on Khintchine, but has not penetrated as fully into the consciousness of philosophers of science [16] as it deserves (although the Ehrenfests consider it carefully in [1]). Statistical Mechanics is defined as the study of the statistical properties of the normal trajectory. By statistical is meant, descriptive statistics, so the notion of probability does not enter into this definition. We, following Wiener, will mean the auto-correlation function $\langle f \rangle_t = \varphi(t)$ of a trajectory. The important thing is to define normal. Jeans [4] defined a normal property of a state (or trajectory) to be a property which is possessed by the overwhelming majority of states in the system, so that, as the number of degrees of freedom increases without bound, the states which do not possess that property possess negligible Liouville measure. A state is then defined as a normal state if it possesses all those normal properties ‘of which it is capable.’ This definition of normal was not given with full logical rigour: the task of this paper is to fix that in an important example.

Some random finite trigonometric sums

Normalise the measure on the surface of constant energy, \mathcal{M}_E , inherited from Liouville measure to be total mass unity. In this paper the only properties we are concerned with are $\varphi(t)$. That is, a normal cell is a sequence of subsets \mathcal{N}_n of $(M_n)_{E_n}$ such that: for every choice of three positive epsilons, we have for n sufficiently large that it has measure $1 - \epsilon$ and the $\varphi(t)$ are within ϵ_1 of each other for $t < 1/\epsilon_2$. The energy level E_n is defined for traditional reasons, and to make the comparison with traditional results convenient, to be that energy level which is most probable according to the Maxwell distribution: it is $\frac{n+1}{kT}$.

Intuitively, this would mean that the limits of the auto-correlation functions (uniform

convergence on compact sets) are the same for points in the same cell.

The existence of a normal cell is the kind of approximate ergodicity analogous to what Khintchine envisioned, (as is, in a very different way, the dispersion theorem proved by Khintchine and Lanford). It is well-known that the limit stochastic process Ford–Kac–Mazur constructed from this sequence is ergodic, since it is the Ornstein–Uhlenbeck process.

It is elementary that in general all for any trigonometric sum

$$p_o(t) = \sum_i a_i \cos(\omega_i t) + b_i \sin(\omega_i t),$$

(here, we may and do assume all $\omega_i > 0$) the auto-correlation function is

$$\varphi(\tau) = \sum_i \frac{1}{2} (|a_i|^2 + |b_i|^2) \cos(\omega_i t).$$

Applying this result to our formula for p_o , we obtain

$$\varphi(\tau) = \sum_k \frac{1}{2} \left(\frac{1}{2N+1} \right)^2 (|\widehat{p}(k)|^2 + |\omega_k \widehat{q}(k)|^2) \cos(\omega_k t).$$

In order to show that normal cells exist, we want to show that the measure of initial conditions in phase space which yield approximately the same $\varphi(\tau)$ tends to unity as $n \rightarrow \infty$. We may regard this as a random trigonometric sum. (Regarding the $p_i(0)$ and the $q_i(0)$ as random variables). We show that the variance of $\varphi(\tau)$ is negligible for large n . In 1866 (see Stroock [15] p. 77) Mehler proved that for $x_1^2 + x_2^2 + \dots x_n^2 = \rho n$ and the uniform surface area measure on the surface of this sphere of total mass one, then as $n \rightarrow \infty$, x_1 tends weakly to a Gaussian random variable with mean zero and standard deviation ρ . The rate of convergence can be controlled explicitly, this is merely a concrete calculation of surface areas on spheres.

It is obvious that the coordinates are uncorrelated, that each x_i is perfectly uncorrelated with the x_j^2 for $j \neq i$, and that the squares are negatively correlated with each other. Hence $Var(x_i^2)$ is approximately two, and $Var(x_i^2 + x_j^2) \leq Var(x_i^2) + Var(x_j^2)$.

We wish to estimate the variances of the $\hat{p}(k)$ and the $\omega_k \hat{q}(k)$.

Because $H(p, q) = H(\hat{p}, \hat{q})$ and, more precisely, because

$$(q_{-N}(0), \dots, q_N(0)) \cdot A \cdot \begin{pmatrix} q_{-N}(0) \\ q_{-N+1}(0) \\ \vdots \\ q_N(0) \end{pmatrix} = (\hat{q}(k), \dots, \hat{q}(k)) (\zeta^{-ml}) \cdot A \cdot (\zeta^{lm}) \begin{pmatrix} \hat{q}(k) \\ \hat{q}(k) \\ \vdots \\ \hat{q}(k) \end{pmatrix}$$

(this is because the matrix $(\zeta^{ml})_{lm}$ is the change of basis matrix that diagonalises A) it follows that

$$(q_{-N}(0), \dots, q_N(0)) \cdot A \cdot \begin{pmatrix} q_{-N}(0) \\ q_{-N+1}(0) \\ \vdots \\ q_N(0) \end{pmatrix} = \sum_k \hat{q}(k)^2 \omega_k^2.$$

But $2H - \sum_k \hat{q}(k)^2 \omega_k^2 = \sum_i p_i(0)^2$. Hence $\sum_k \hat{q}(k)^2 \omega_k^2$ and $\sum_i p_i(0)^2$ are perfectly anti-correlated and hence have equal variances. But this last is bounded by $2(2N+1)$. Now the ω_k are all real and positive in the applications we have in mind for later. Also, since the matrix is symmetric, $\omega_{-k} = \omega_k$. Hence, obviously, we may arrange that the $\hat{q}(k)$ are real without altering the auto-correlation function (since $\hat{q}(k) = \overline{\hat{q}(-k)}$). Hence

$$Var \frac{1}{2N+1} \frac{\frac{1}{2}}{2N+1} \sum_k |\omega_k \hat{q}(k)|^2 \cos(\omega_k \tau) < \frac{1}{(2N+1)^2} \frac{\frac{1}{4}}{(2N+1)^2} 2(2N+1).$$

Since, by definition, $\hat{p}(k) = \sum_i \zeta^{-ik} p_i(0)$, we have that $\frac{\hat{p}(k)}{\sqrt{2N+1}}$ is related to $p_i(0)$ by a unitary transformation, the sum of squares of the moduli of the coordinates does not change. Then neither does their variance. Hence the variance of $\frac{\sum \frac{1}{2} \hat{p}(k)}{(2N+1)^2}$ is bounded by $\frac{2(2N+1) \frac{1}{4}}{(2N+1)^2} = \frac{\frac{1}{2}}{(2N+1)}$.

This very weak ergodicity does not depend on any properties of the ω_k except that the matrix A is cyclic and symmetric. It holds even when the dynamics does not tend towards Brownian motion (which only happens for a very specific choice of ω_k).

We wish to relate Wiener's time auto-correlation function, a deterministic concept, to the phase auto-correlation function, at least when the trajectory is normal. Note first that as usual, the Maxwell distribution 'bunches up' for very large n around the most probable energy value, and thus the average taken with respect to the Maxwell distribution over the entire, unbounded, phase space is the same as the average over one energy level ellipsoid, with respect to the uniform distribution. An elementary part of what Khintchine proved [9] p. 68, is

Theorem (Khintchine): Let $d\mu_E$ be the normalised measure on the constant energy shell Ω_E of \mathcal{M}_n with energy level $H = E$ inherited from Liouville measure. Then $d\mu_E$ is invariant under the flow. Let $\varphi^\omega(\tau)$ be Wiener's time-autocorrelation function for a given initial condition $\omega \in \Omega_E$. Then the expectation of the time-autocorrelation function is equal to the phase auto-correlation function, *i.e.*,

$$\int_{\Omega_E} \varphi^\omega(\tau) d\mu(\omega) = \int_{\Omega_E} f_t(\omega) f_{t+\tau}(\omega) d\mu(\omega).$$

Time auto-correlation functions of the Ornstein–Uhlenbeck process

We now specify a precise dynamics by explicitly choosing the ω_k . Let $\omega_s = \tan \frac{\pi s}{n+1}$.

Ford–Kac–Mazur calculate the phase auto-correlation of each finite stage and pass to the limit (with a cut-off renormalisation) obtaining the usual auto-correlation function of the Ornstein-Uhlenbeck process, $\pi e^{-|\tau|}$.

It is part of what they prove that for any compact set $K = \{\tau \in [0, K]\}$, there exists

an N so large that

$$\int_{\Omega_E} p_o(0)(p_o)_\tau d\mu(\omega)$$

approximates to $\pi e^{-|\tau|}$ to any desired accuracy.

Their method of proof was relatively elementary and can be included here. A simple trick changes this into a standard cosine transform which can be looked up in any table of integrals. Let $u = \tan \theta$. Then $\theta = \arctan u$. This integral is, then, equal to one of the Riemann sums for the improper integral

$$\int_{-\infty}^{\infty} \cos(\tau u) \frac{du}{u^2 + 1}$$

This, as aimed at, is the cosine transform of the bump function $\frac{1}{u^2+1}$. It is equal to $\pi e^{-\tau}$ when τ is positive, but it is symmetric since cosine is an even function, so it is equal to $\pi e^{-|\tau|}$. There is no problem with convergence in this calculation as we let $\epsilon \rightarrow 0$; the improper integral is very nicely behaved.

Theorem: Suppose given $\epsilon, \delta > 0$ and K . Then there exists an N so large that the measure of the set of trajectories such that $|\varphi(\tau) - \pi e^{-|\tau|}| > \epsilon$ for all $\tau \in K$ is less than δ . (This implies that a normal cell exists, the cell of all trajectories such that on K , their auto-correlation functions are within ϵ of the phase auto-correlation function $\langle f_o f_\tau \rangle$.)

Proof. Since the variance of $\varphi(\tau)$ for any fixed τ is less than $\frac{1}{(2N+1)}$, Tchebycheff's inequality gives us that the measure of the set of trajectories such that $\varphi(\tau)$ differs by more than ϵ from $\langle f_o f_\tau \rangle$, which is its expected value, is less than $\frac{1}{\epsilon n}$. If we cover the interval K with a uniform mesh of width Δx then there are $\frac{K}{\Delta x}$ values of τ_i in this mesh. If we treat the measure of the set of trajectories yielding a deviation of $\varphi(\tau_i)$ by more than ϵ from its expectation for each i as independent events, which they are not, then the

measure of the set of initial conditions such that even one of these violations will occur is less than $\frac{K}{n\epsilon\Delta x}$.

As n grows, the difference between the quasi-periodic expectation of $\varphi(\tau)$ and its limit the exponential decay $\pi e^{-|\tau|}$, may be arranged to be less than ϵ on any compact set, especially K . And this latter function clearly satisfies $|\Delta y| < \Delta x$ over this or any other mesh. Hence, if τ varies over a region of width Δx , $\varphi(\tau)$ varies by less than $2\epsilon + \Delta x$. We may take $\Delta x = .5\epsilon$. and $n > K\frac{2}{\epsilon^2}$ to get a 4ϵ proof. *Q.E.D.*

This method yields an essentially independent proof of the results of Ford–Kac–Mazur and their generalisations. It does not seem as if the usual methods of large-sample theory of time series, or of Khintchine, Ruelle, or Lanford, or the usual limit theorems of probability theory, can be used to obtain this or similar results. I would like to conjecture that normal cells in this sense exist for a much wider class of sequences of dynamical systems.

Now for all practical purposes, a stochastic process can be replaced by its auto-correlation function. In fact, a Gaussian stationary stochastic process is determined up to equivalence by the phase auto-correlation function. As n increases without bound, the time auto-correlation functions of normal trajectories approaches a limiting function. We may define a stochastic process by the requirement that its phase auto-correlation function be this limiting function. Thus we have defined the thermodynamic limit of this sequence of dynamical systems as a stochastic process. The sample space of this process has nothing to do with any Hamiltonian dynamical system or Liouville measure. There does not seem to me to be any point in trying to define a new class of dynamical system, with an infinite number of degrees of freedom, which would be the limit object here: this would go against Lévy’s philosophy, and would be subject to the use of Occam’s razor.

In [5] and [6], I have shown how a quantum analogue at negative temperature, which is much simpler than the classical case, has many of the same features as the model of this paper. It would be important to generalise the results of this paper to a negative temperature heat bath around the mote. Schwinger in [14] has treated the case of quantum negative temperature Brownian motion, claimed it acts as an amplifier (which is understandable, since it is done in [5]), and claimed that it amplifies quantum motion to the classical level. The derivation lacks rigour and uses the usual imprecise notions of probability. This is an important topic for the future.

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